Estimation of Parameter Values

Nutrition Models Workshop

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Outline

- Nutrition models are VERY diverse
 - Combination of empirical, mechanistic, dynamic and static models
 - Regression, linear and nonlinear mixed models, differential equations

- Today: Main approaches for estimating parameters in a variety of models
 - Some mathematical description
 - Idea is for you to understand the reasoning and challenges of different approaches
- One exercise/demonstration in the end
 - Fit model with two approaches

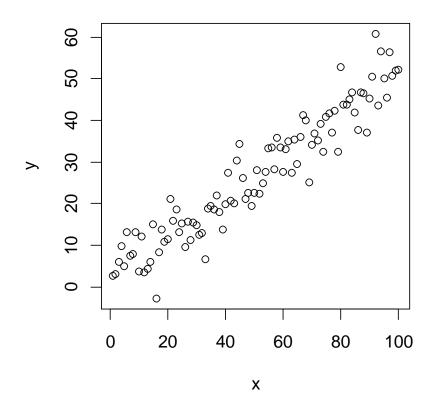
Introduction

- Different types of models have been used for nutrition modeling
 - Compartmental, regression, meta-analysis, nonlinear mixed models,

- One feature is common to almost all these models
 - Parameters are needed to describe the system
 - Quantify relationship between variables

Introduction

• Simple example: linear regression



$$Y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$$

- Y_i is the response variable for the *i*th observation
- x_i is the predictor variable in the *i*th observation
- β_0 is the intercept
- β_1 is the slope
- ε_i is the error, $E[\varepsilon_i] = 0$, $Var[\varepsilon_i] = \sigma^2$ and ε_i are independent
- *i* = 1, ..., *n*

In matrix notation: $\mathbf{y} = \mathbf{X}\mathbf{\beta} + \mathbf{\epsilon}$

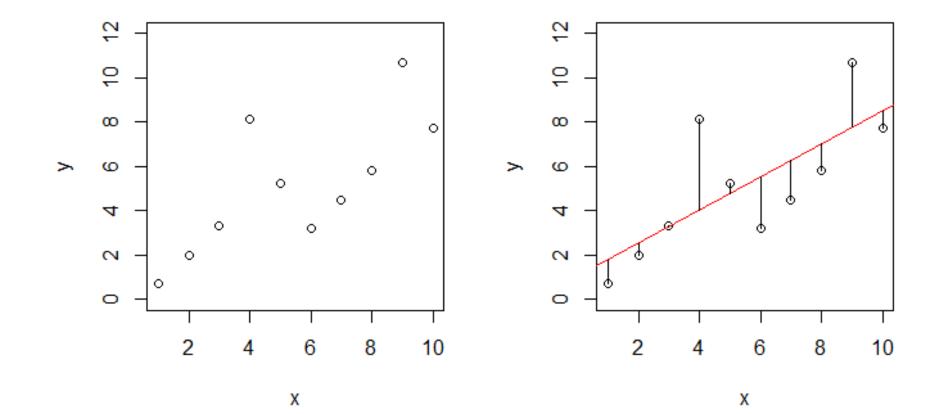
$$\begin{bmatrix} Y_1 \\ \vdots \\ Y_n \end{bmatrix} = \begin{bmatrix} 1 & x_1 \\ \vdots & \vdots \\ 1 & x_n \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \end{bmatrix} + \begin{bmatrix} \varepsilon_1 \\ \vdots \\ \varepsilon_n \end{bmatrix}$$

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Introduction

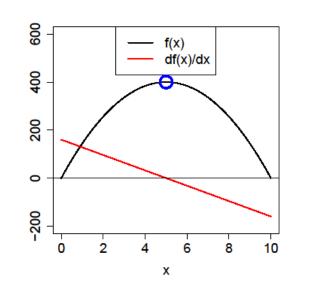
- In practice, parameter true values are unknown
 - Estimators from a sample

- Parameters have to be optimal in some sense
 - Least square estimators minimize squared errors
 - Maximum likelihood estimators maximize the likelihood function



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- The least squares estimators minimize the square errors: $Q = \sum_{i=1}^{n} [Y_i (\beta_0 + \beta_1 x_i)]^2$
- How do we get them???
- We can find points of minimum and maximum of a function using derivatives.



For example for $f(x) = 160x - 16x^2$

Set derivative to zero and "solve" for *x*:

$$160 - 32x = 0$$
$$x = 5$$

Second derivative test: -32

$$\frac{\partial Q}{\partial \beta_0} = -2\sum_{i=1}^n (Y_i - \beta_0 - \beta_1 x_i)$$
$$\frac{\partial Q}{\partial \beta_0} = -2\sum_{i=1}^n x_i (Y_i - \beta_0 - \beta_1 x_i)$$

Setting these partial derivatives to zero, we construct the normal equations

$$\sum_{i=1}^{n} Y_{i} = nb_{0} + b_{1} \sum_{i=1}^{n} X_{i}$$
$$\sum_{i=1}^{n} X_{i}Y_{i} = b_{0} \sum_{i=1}^{n} X_{i} + b_{1} \sum_{i=1}^{n} X_{i}^{2}$$

• The least square estimators are the solutions to the normal equations

$$b_{1} = \sum_{i=1}^{n} \left(x_{i} - \overline{x} \right) \left(Y_{i} - \overline{Y} \right) / \sum_{i=1}^{n} \left(x_{i} - \overline{x} \right)^{2}$$
$$b_{0} = \overline{Y} - b_{1} \overline{x}$$

• The concept extends to multiple regression

$$Q = \sum_{i=1}^{n} \left[Y_i - \left(\beta_0 + \beta_1 x_{i1} + \beta_{p-1} x_{i,p-1} \right) \right]^2$$

• General form of the least squares estimators: $\mathbf{b} = (\mathbf{X}^{\mathrm{T}}\mathbf{X})^{-1}\mathbf{X}^{\mathrm{T}}\mathbf{Y}$

- Estimates of the uncertainty associated with these parameters
- Estimator of the error's variance

$$MSE = \frac{1}{n-p} \sum_{i=1}^{n} \left[Y_i - \left(b_0 + b_1 x_{i1} + \dots + b_{p-1} x_{i,p-1} \right) \right]^2$$

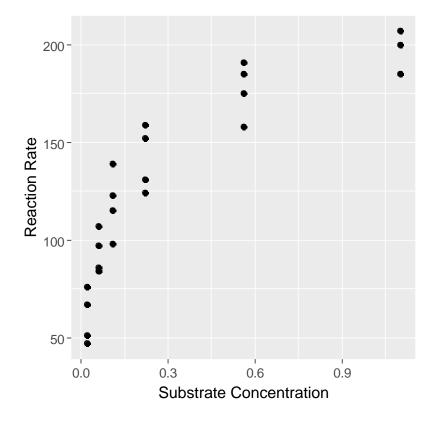
• Estimated variance-covariance matrix of the parameters $MSE(\mathbf{X}^{T}\mathbf{X})^{-1}$

Nonlinear Models

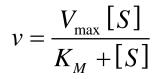
- So far, we can estimate parameters in linear models
- Many phenomena in biology are nonlinear
 - For example, reaction velocity vs. substrate concentration in an enzymatic reaction
 - Before we start with nonlinear models, let's clarify

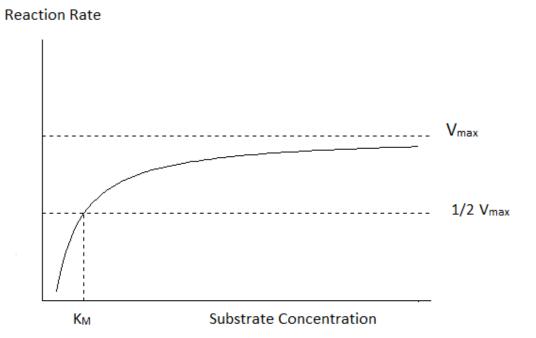
$$Y_{i} = \frac{\theta_{1}}{1 + \theta_{2} \exp(\theta_{3} x_{i})} + \varepsilon_{i} \quad \text{is a nonlinear model}$$
$$Y_{i} = \beta_{0} + \beta_{1} x_{i} + \beta_{2} x_{i}^{2} + \varepsilon_{i} \quad \text{is a linear model}$$

Nonlinear Models



Michaelis-Menten Kinetics





Nonlinear Regression

 $Y_i = f\left(x_i, \boldsymbol{\theta}\right) + \mathcal{E}_i$

f is the nonlinear function describing the relationship between Y and x

Michaelis-Menten example: $y_i = \frac{V_{\max} x_i}{K_M + x_i} + \varepsilon_i$

$$\frac{Y_{\max} x_i}{M_{\max} + x_i} + \mathcal{E}_i \qquad \Longrightarrow f\left(x_i, \mathbf{\theta}\right) = \frac{V_{\max} x_i}{K_M + x_i}$$

- y_i is the reaction rate for the *i*th observation
- x_i is the associated substrate concentration

 $\mathbf{\theta} = \left(V_{\max}, K_{M}\right)^{\mathrm{T}}$ are the parameters to be estimated

- ε_i is the error, $E[\varepsilon_i] = 0$, $Var[\varepsilon_i] = \sigma^2$ and independent
- *i* = 1, ..., *n*

• For the simple linear regression model, least squares minimize

$$Q = \sum_{i=1}^{n} \left[Y_{i} - (\beta_{0} + \beta_{1} x_{i}) \right]^{2}$$

• For the nonlinear regression, the idea is the same: minimize

$$Q = \sum_{i=1}^{n} \left[Y_i - f(x_i) \right]^2$$

 Solution to the normal equations are often difficult to obtain analytically

- Numerical Algorithms
 - For example, Gauss-Newton
 - Require initial values to initialize numerical procedures

Gauss-Newton

- Default in PROC NLIN and ${\tt nls}$ ()
- Approximate the nonlinear model with linear terms
- Taylor series expansion and least squares as for linear regression
- Denote the least squares estimates g and the initial values $g^{(0)} = (g_0^{(0)}, g_1^{(0)}, \dots, g_{p-1}^{(0)})$
- Approximation around starting values:

$$f(x_i, \mathbf{\theta}) \approx f(x_i, \mathbf{g}^{(0)}) + \sum_{k=0}^{p-1} \left[\frac{\partial f(x_i, \mathbf{\theta})}{\partial \theta_k} \right]_{\mathbf{\theta} = \mathbf{g}^{(0)}} \left(\theta_k - g_k^{(0)} \right)$$

Gauss-Newton

Model approximation

$$Y_{i} - f\left(x_{i}, \mathbf{g}^{(0)}\right) \approx \sum_{k=0}^{p-1} \left[\frac{\partial f\left(x_{i}, \mathbf{\theta}\right)}{\partial \theta_{k}}\right]_{\mathbf{\theta}=\mathbf{g}^{(0)}} \left(\theta_{k} - g\right)_{k}^{(0)} + \varepsilon_{i}$$

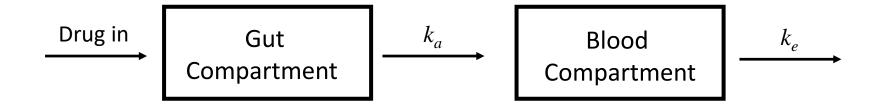
- It is a linear model!
- Estimate parameters by least squares: $\mathbf{b}^{(0)} = \left(\mathbf{D}^{(0)T}\mathbf{D}^{(0)}\right)^{-1}\mathbf{D}^{(0)T}\mathbf{Y}^{(0)}$
- Update: $g^{(1)} = g^{(0)} + b^{(0)}$

Gauss-Newton

- Evaluation criteria: $SSE^{(0)} = \sum_{i=1}^{n} (Y_i f_i^{(0)})^2$
- Start the process again with $\mathbf{g}^{(1)}$ as the initial values
- Repeat procedure until $SSE^{(s+1)} SSE^{(s)}$ is negligible
- Estimate of error's variance: $MSE = \sum_{i=1}^{n} \left[Y_i f(x_i, \mathbf{g}) \right]^2 / n p$
- Other methods available, e.g. Nelder-Mead and Marquardt

Compartmental Models

- Traditionally used in nutritional modeling
 - Roots on pharmacokinetics and differential calculus



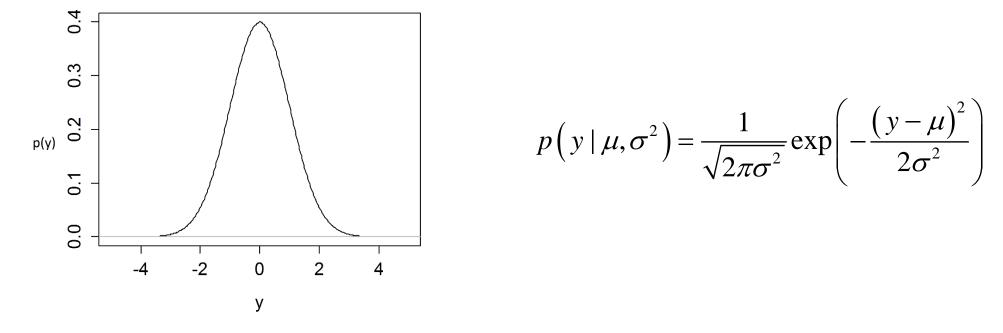
Compartmental Models

- Functional forms described in terms of differential equations
 - Instead of the "integrated form"
- Strategy for parameter estimation
 - Expected mean represented by a compartmental model f
 - If *f* cannot be obtained analytically, it has to be solved numerically
 - Euler, Runge-Kutta4, Isoda
 - Can use nonlinear least squares but have to numerically solve f at iteration
 - Modern software estimate using maximum likelihood

Maximum Likelihood Estimation

- Another strategy for parameter estimation
- For regression models with independent $\varepsilon_i \sim N(0, \sigma^2)$, estimators coincide with least squares estimators
- Estimators maximize the likelihood function
 - Parameter values that are in best agreement with the data

Maximum Likelihood Estimation



- $p(y | \mu \sigma^2)$ is the density function: How likely y is at each value
- The likelihood function is: $L(\mu, \sigma^2 | y_1, ..., y_n) = p(y_1 | \mu, \sigma^2) \times ... \times p(y_n | \mu, \sigma^2)$
 - "How likely the whole data is with that set of parameters values"
 - MLE: "maximize the likelihood of getting the observed data"

Maximum Likelihood Estimation

• Linear Regression Example: $Y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$

$$L(\boldsymbol{\beta}_0, \boldsymbol{\beta}_1, \sigma^2 \mid y_1, \dots, y_n) = \left(\frac{1}{\sqrt{2\pi\sigma^2}}\right)^n \exp\left[-\frac{1}{2\sigma^2}\sum_i \left(y_i - \boldsymbol{\beta}_0 - \boldsymbol{\beta}_1 x_i\right)^2\right]$$

• It is easier to work with the log-likelihood

$$\log L(\beta_0, \beta_1, \sigma^2 | y_1, \dots, y_n) = -\frac{n}{2} \log (2\pi) - \frac{n}{2} \log (\sigma^2) - \frac{1}{2\sigma^2} \sum_i (y_i - \beta_0 - \beta_1 x_i)^2$$

- To find parameters that maximize the likelihood we...
 - Take the derivative with respect to each parameter and set to zero. Also need second derivative test

Mixed Models

- Modern mixed modeling relies heavily on likelihood methods
- Extension of (non)linear models with both fixed and random effects
- Probably the "type" of model you need to analyze your data or construct your nutrition model
- There's a whole workshop on mixed models in this meeting

Nonlinear Mixed Models

- Nonlinear functional forms
 - Michaelis-Menten, logistic, exponential, Gompertz, ...

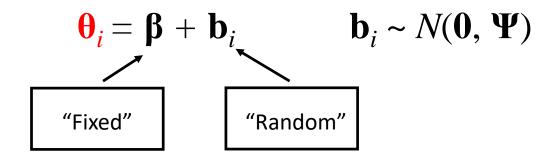
• Random effects that "enter the model" nonlinearly

- Allow you to model nonlinear clustered, logitudinal data
 - Records from the same animal, treatment means from the same study

Nonlinear Mixed Models

$$y_{ij} = f\left(x_{ij}, \mathbf{\theta}_i\right) + \mathcal{E}_{ij}$$

- y_{ij} is the *j*th record on the ith "subject" or cluster
- x_{ii} is the associated predictor variable
- θ_i is the vector of subject specific parameters



• ε_{ij} is the random error ~ $N(0, \sigma^2)$

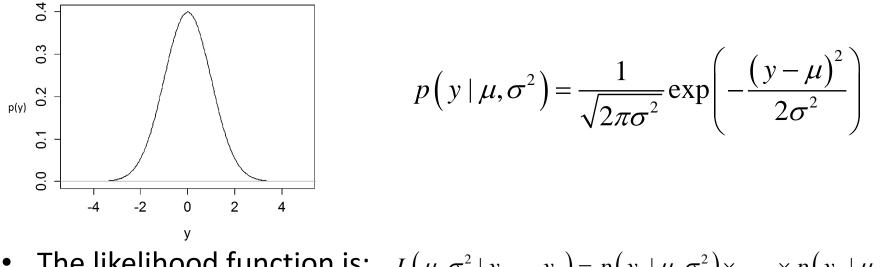
Nonlinear MM Maximum Likelihood

- There is more than one source of variability
 - Between subjects and within subjects

- To represent the generative process of the data we need to take both into account
 - Joint density of the response and the random effects

Maximum Likelihood

• For the linear regression



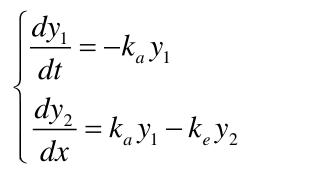
• The likelihood function is: $L(\mu, \sigma^2 | y_1, ..., y_n) = p(y_1 | \mu, \sigma^2) \times ... \times p(y_n | \mu, \sigma^2)$

• For the nonlinear mixed model, we need to compute the marginal density of the responses: $p(\mathbf{y} | \boldsymbol{\beta}, \sigma^2, \Psi) = \int p(\mathbf{y} | \mathbf{b}, \boldsymbol{\beta}, \sigma^2) p(\mathbf{b} | \Psi) d\mathbf{b}$

Maximum Likelihood

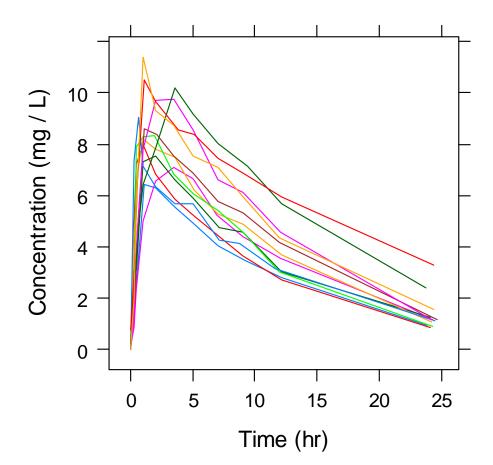
- Bad news from a estimation perspective
- The likelihood function to estimate the parameters requires integrating the joint density with respect to the random effects
- The integral often does not have a closed form expression
- Approximation of the likelihood function

• Let's go back to the compartmental model



- y_1 is the amount of drug in the gut compartment
- y_2 is the amount of drug in the blood compartment
- k_a is the absorption rate (1/hr)
- k_e is the elimination rate (1/hr)

- Data from Davidian and Giltinan (1995)
- 12 subjects received a single oral dose of theophylline
 - Anti-asthmatic drug
 - Single oral dose at time zero
- Measurements of blood concentrations of drug at 11 time points over a 25 hour period



- We only observe data from the second compartment
- The differential equation for the second compartment can actually be solved analytically
- We will estimate the parameters in two different ways
 - Analytical solutions: nonlinear mixed model
 - Numerical solutions for differential equations in a mixed model framework

• The analytical solution to the second differential equation is

$$C(t) = \frac{Dose \, k_e k_a}{Cl \left(k_a - k_e\right)} \left[\exp\left(-k_e t\right) - \exp\left(-k_a t\right) \right]$$

• As a nonlinear mixed model

$$C_{ij} = \frac{Dose_i k_e k_{a,i}}{Cl_i (k_{a,i} - k_e)} \Big[\exp(-k_e t_{ij}) - \exp(-k_{a,i} t_{ij}) \Big] + \varepsilon_{ij}$$

$$\varepsilon_{ij} \sim N(0, \sigma^2)$$
 and $\begin{bmatrix} Cl_i \\ k_{a,i} \end{bmatrix} \sim N\left(\begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} \sigma_{Cl}^2 & 0 \\ 0 & \sigma_{k_a}^2 \end{bmatrix}\right)$

- All three parameters must be positive
- Reparameterize model with parameters on a log scale

$$C(t) = \frac{Dose \exp\left(lk_e + lk_a - lCl\right)}{\exp\left(lk_a\right) - \exp\left(lk_e\right)} \left\{ \exp\left[-\exp\left(lk_e\right)t\right] - \exp\left[-\exp\left(lk_a\right)t\right] \right\}$$

where

$$lk_e = \log(k_e)$$
, $lk_a = \log(k_a)$ and $lCl = \log(Cl)$

- First method
 - Fit nonlinear mixed model in R

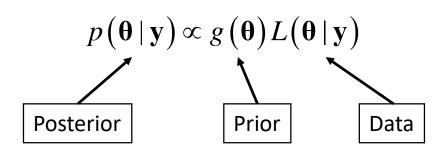
- Second method
 - Nonlinear mixed model but solve differential equations numerically: lsoda solver

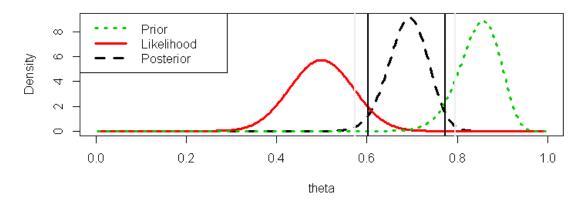
$$\begin{cases} \frac{dy_1}{dt} = -k_a y_1 \\ \frac{dy_2}{dx} = k_a y_1 - k_e y_2 \\ y_1(0) = Dose \\ y_2(0) = 0 \end{cases}$$

• Observation Equation: $C(t) = \frac{y_2(t)}{Cl}k_e$

Bayesian Inference

- Combines prior information with new data: update of knowledge
- All parameters are treated as random variables
 - Prior distributions for parameters
 - Inference is based on the posterior distribution
 - Bayes theorem





From: https://www.r-bloggers.com/the-beta-prior-likelihood-and-posterior/

Prior: beta(52.22,9.52); Data: B(50,25); Posterior: beta(77.22,34.52)

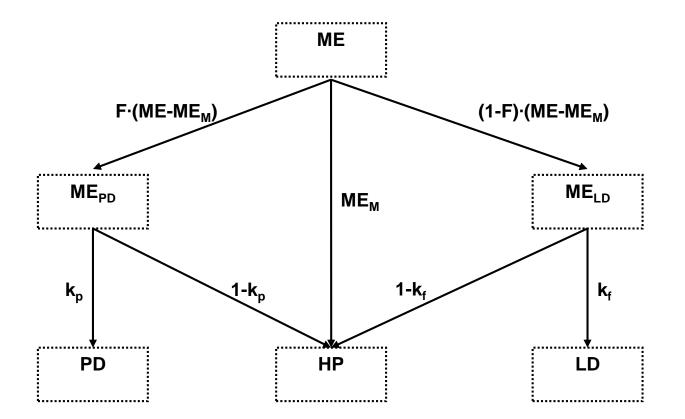
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Bayesian Inference

- Inference based on posterior
 - Combines prior information with the observed data
 - Particularly suited for models built with many parameters that good biological knowledge is available
 - Many freely software available
 - Including with differential equations "solvers"
 - We don't have to have a known or tractable posterior: Markov Chain Monte Carlo (MCMC)

Example of a Multivariate Nonlinear Model

• From Strathe et al. (2012). J. Agri. Sci. 150:764-774



Example for a Multivariate Nonlinear Model

$$PD = \frac{PD_{max}}{BW_{PDmax}} BW \cdot \log\left(\frac{BW_{PDmax} \cdot \exp(1)}{BW}\right)$$
$$LD = k_f \left(ME - a \cdot BW^b - PD / k_p\right)$$

Multivariate Model:

Priors from literature

partial etticiencies k_p and k_f

k _f	k _p	Reference
0.72–0.88	0.52-0.63	Strathe <i>et al.</i> (2010 <i>a</i>)*
0.75	0.56	van Milgen <i>et al.</i> (2000)*
0.77–0.82	0.58-0.60	van Milgen & Noblet (1999)*
0.84	0.62	Noblet <i>et al.</i> (1999)†
0.76	0.54	NRC (1998)†
0.60	0.52	Tess <i>et al.</i> (1984)†
0.74	0.56	ARC (1981)†

* Partial efficiencies derived from multivariate modelling $k_p \sim apple (0.80, 0.10^2)$ and $k_f \sim N(0.80, 0.10^2)$

Credible Intervals for the Posterior

Ь	0.60 (0.56-0.66)	0.61 (0-
k_p	0.59 (0.53-0.65)	0.58 (0-
k_{f}	0.78 (0.72-0.85)	0.76 (0

From Strathe et al. (2012). J. Agri. Sci. 150:764-774